

# Markov Renewal Methods in Restart Problems in Complex Systems

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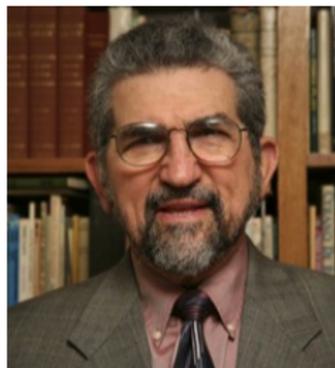


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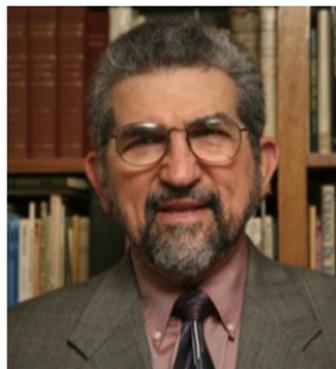
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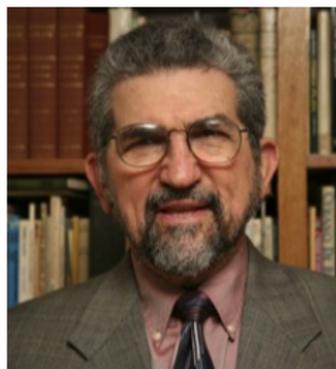
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- **Call centers** — 'customer service' by telephone.  
Failures due to broken connection etc.

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- $\mathbb{E}X = ??$  Easy
- $\mathbb{P}(X > x) = ??$  Main problem here

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SA–Fiorini-Lipsky-Rolski-Sheahan

*Mathematics of Operations Research* 2008

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Cramér-Lundberg asymptotics: geometric sums, renewal equation

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Poisson failures,  $L$  gamma, shape  $\alpha$ :  $\mathbb{P}(X > x) \sim C \frac{\log x^{\alpha-1}}{x^\beta}$

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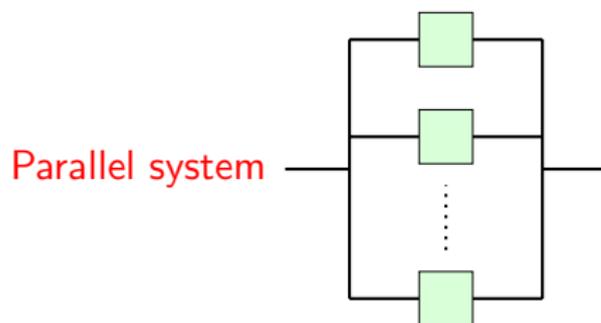
$N > n \iff U_1 < L, \dots, U_n < L \iff L = \max(L, U_1, \dots, U_n)$

$$\mathbb{P}(N > n) = \frac{1}{n+1}$$

# Complex Systems

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## Classical reliability theory

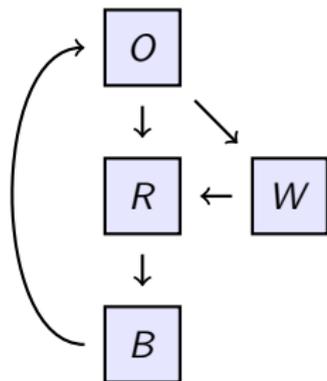


$k$ -out-of- $n$

Repair; cold/warm standby; . . .

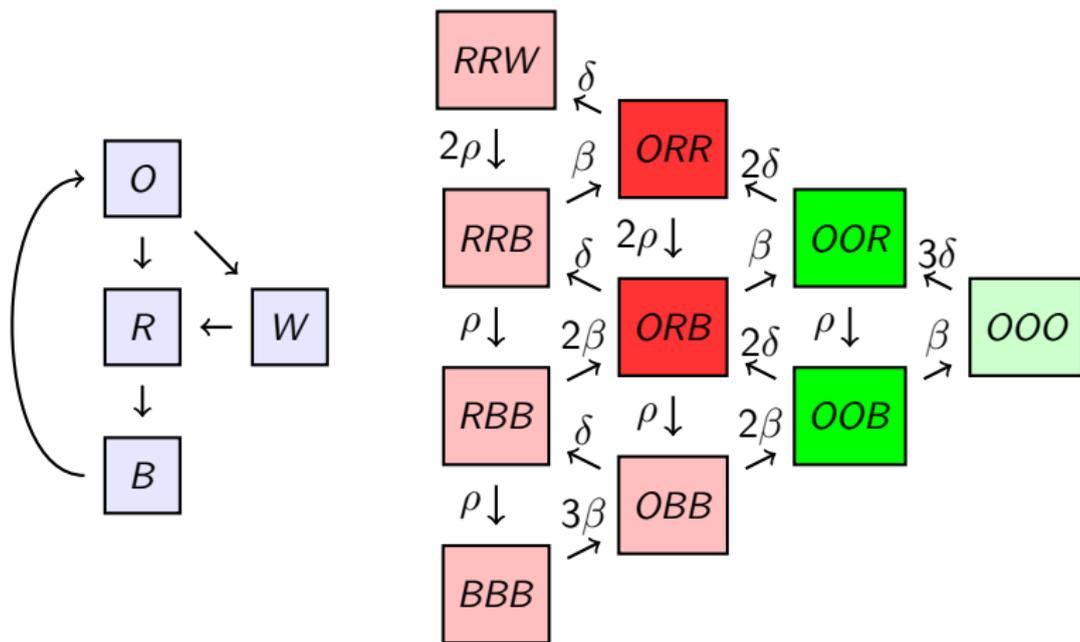
2 - out - of - 3 processors, 2 repairmen, warm standby

Processor phases  $\left\{ \begin{array}{l} \text{Operating} - \delta \\ \text{Repair} - \rho \\ \text{Booting} - \beta \end{array} \right. \quad \text{Waiting}$

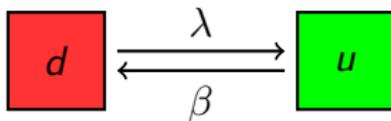


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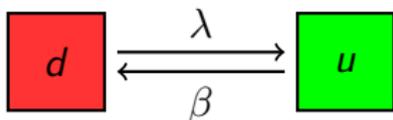


# Markov renewal equation. I



$$L \equiv \ell$$

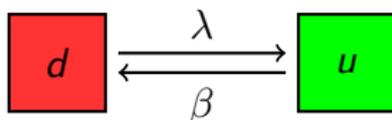
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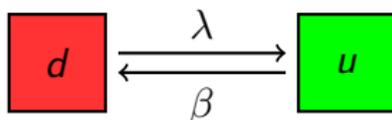


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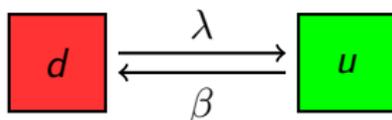
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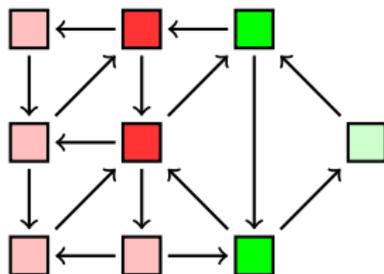
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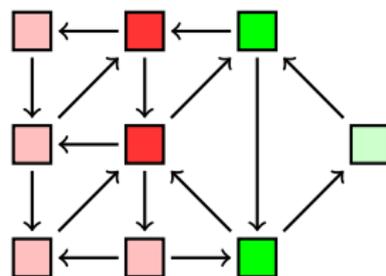
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## Theorem

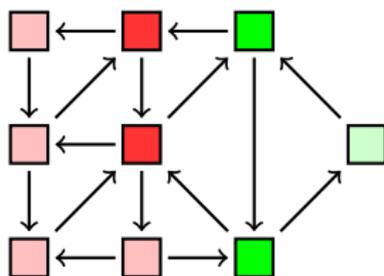
$\mathbb{P}_u(X > x) \sim C e^{-\gamma x}$  where  $\gamma > 0$  solves

$$1 = \frac{\lambda \beta}{(\lambda - \gamma)(\beta - \gamma)} [e^{(\gamma - \beta)\ell} - 1] \quad \text{and } C = \dots$$





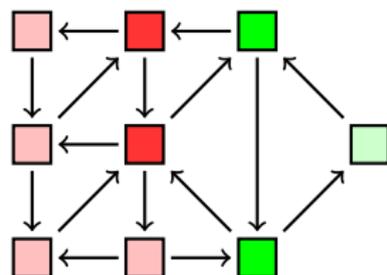
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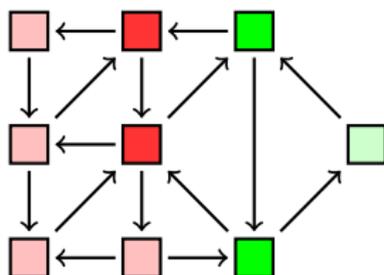


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Markov renewal state space  $\mathcal{E} = \mathcal{U} \cup \mathcal{D}$

Sojourn time  $T_i$  in  $i \in \mathcal{E}$  depends on full generator matrix

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## Theorem

Denote by  $\mathbf{R}(\alpha)$  the  $\mathcal{E} \times \mathcal{E}$  matrix with entries

$$r_{du}(\alpha) = \mathbb{E}_d[e^{\alpha T_d}; \xi_1 = u], \quad d \in \mathcal{D}, u \in \mathcal{U},$$

$$r_{ud}(\alpha) = \mathbb{E}_u[e^{\alpha T_u}; \ell \geq T_d, \xi_1 = d], \quad u \in \mathcal{U}, d \in \mathcal{D},$$

all other  $r_{ij}(\alpha) = 0$ . Assume there exists  $\gamma = \gamma(\ell)$  such that  $\mathbf{R}(\gamma)$  is irreducible with  $\text{spr}(\mathbf{R}) = 1$ . Then  $\mathbb{P}_i(X > x) \sim C_i e^{-\gamma x}$ ,  $x \rightarrow \infty$

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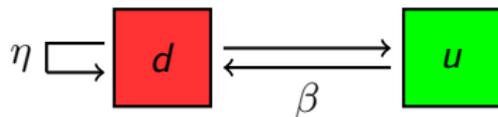
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Markov renewal case: SA - Thøgersen 2015

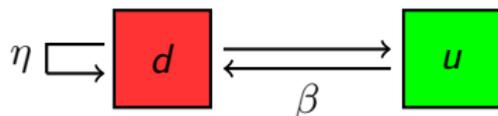
## Heavy-tailed example

Ideal repair time  $R$  (random); rate  $\eta$  failures of repair



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Actual repair time: vanilla Restart

If  $R$  is Gamma:  $\mathbb{P}(X > x) \sim C \frac{\log^{\alpha-1} x}{x^\mu}$

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Know  $\mathbb{P}_i(X > x | L = \ell) \sim C_i e^{-\gamma(\ell)x}$  (with light  $T_d$  tails)  
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## Corollary

*If the task length  $L$  has unbounded support, the distribution of the total task time  $X$  is heavy-tailed in the sense that  $e^{\delta x} \mathbb{P}(X > x) \rightarrow \infty$  for all  $\delta > 0$ .*

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More precise asymptotics?

Needs asymptotics of  $\gamma(\ell)$

## Theorem

Assume that for some function  $\varphi(\ell)$  it holds that

$$\mathbb{P}(T_u > \ell, \xi_1 = d) \sim k_{ud}\varphi(\ell)$$

as  $\ell \rightarrow \infty$  for some set of constants such that  $k_{ud} > 0$  for at least one pair  $u \in \mathcal{U}, d \in \mathcal{D}$ . Then

$$\gamma(\ell) \sim \mu \varphi(\ell) \quad \text{as } \ell \rightarrow \infty, \quad \text{where } \mu = \frac{\sum_{u \in \mathcal{U}, d \in \mathcal{D}} \pi_u k_{ud}}{\sum_{i \in \mathcal{U} \cup \mathcal{D}} \pi_i \mathbb{E}_i T_i}$$

and  $\pi = (\pi_i)_{i \in \mathcal{U} \cup \mathcal{D}}$  is the stationary distribution of the Markov chain  $\xi$ , that is, the invariant probability vector for the matrix  $\mathbf{P} = \mathbf{R}(0, \infty)$ .

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1) **Perturbation theory**  $\mathbf{R}(0, \infty) = \mathbf{P}$   
 $\text{spr}(\mathbf{R}(\gamma, \ell)) = \text{spr}(\mathbf{P}) + ?? = 1 + ??$

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 $\text{spr}(\mathbf{R}(\gamma, \ell)) = \text{spr}(\mathbf{P}) + ?? = 1 + ??$

2) Implicit function theorem

## Proof ??

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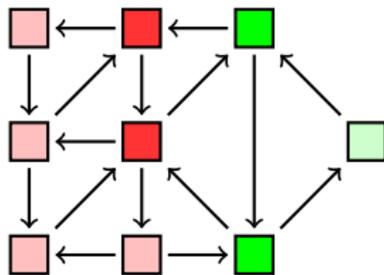
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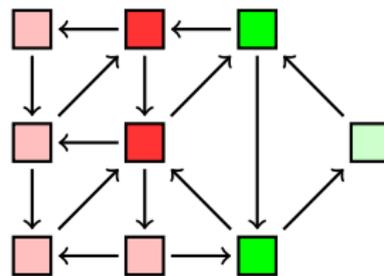
3) Bare-hand (but Perron-Frobenius theory key tool)

## Back to Markov set-up

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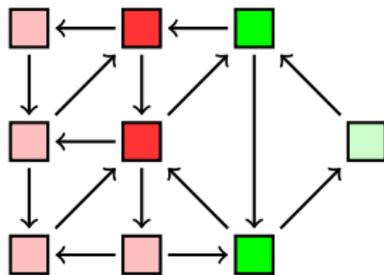


## Back to Markov set-up



Markov renewal up states  $\mathcal{U}$ : two dark green  
Markov renewal down states  $\mathcal{D}$ : two dark red

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Markov renewal down states  $\mathcal{D}$ : two dark red

Markov up states  $\mathcal{U}^*$ : all three green

Markov down states  $\mathcal{D}^*$ : all six red



## Theorem

Assume that  $\gamma = \gamma(\ell)$  makes the spectral radius of the matrix  $\mathbf{R}(\gamma, \ell)$  equal to 1, where  $\mathbf{R}(\gamma, \ell)$  is the matrix

	$\mathcal{U}$	$\mathcal{D}$
$\mathcal{U}$	0	$\hat{F}_{ud}[\gamma]$
$\mathcal{D}$	$\hat{F}_{du}[\gamma]$	0

$$\hat{F}_{ud}[\gamma] = \int_0^\ell e^{\gamma t} F_{ud}(dt)$$

$$\hat{F}_{du}[\gamma] = \int_0^\infty e^{\gamma t} F_{du}(dt)$$

Then  $\gamma(\ell) \sim \mu e^{-\delta \ell}$  as  $\ell \rightarrow \infty$ , where  $-\delta$  is the largest eigenvalue of  $\mathbf{Q}_{\mathcal{U}^* \mathcal{U}^*}$  in the block-partitioning

	$\mathcal{U}^*$	$\mathcal{D}^*$
$\mathcal{U}^*$	$\mathbf{Q}_{\mathcal{U}^* \mathcal{U}^*}$	$\mathbf{Q}_{\mathcal{U}^* \mathcal{D}^*}$
$\mathcal{D}^*$	$\mathbf{Q}_{\mathcal{D}^* \mathcal{U}^*}$	$\mathbf{Q}_{\mathcal{D}^* \mathcal{D}^*}$

of the full generator  $\mathbf{Q}$  and  $\mu$  involves  $\mathbf{Q}_{\mathcal{U}^* \mathcal{U}^*}^{-1}$ ,  $\mathbf{Q}_{\mathcal{D}^* \mathcal{D}^*}^{-1}$  and further Perron-Frobenius characteristics of  $\mathbf{Q}$ .

## General approach for random task time $L$

$f(\ell)$  density of  $\ell$

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Most often OK; now purely analytical problem.

### Corollary

Assume failures are Poisson( $\delta$ ) (or  $\varphi(\ell) = e^{-\delta\ell}$ ) and that  $F$  is gamma-like in the sense that  $f(\ell) \sim c_F \ell^{\alpha-1} e^{-\lambda\ell}$ ,  $\ell \rightarrow \infty$ . Then

$$\mathbb{P}_i(X > x) \sim \frac{C_i^* \Gamma(\lambda/\delta) \log^{\alpha-1} x}{\delta^{\alpha+\lambda/\delta} x^{\lambda/\delta}} \text{ as } x \rightarrow \infty.$$

4×4 table of examples of rough  $\mathbb{P}(X > x)$  asymptotics  
 each of  $f(\ell), \varphi(\ell)$  LT Weibull; exponential; HT Weibull; power

$f(\ell)$	$e^{-\ell^2}$	$e^{-\ell}$	$e^{-\ell^{1/2}}$	$\frac{1}{\ell^\alpha}$
$\varphi(\ell)$				
$e^{-\ell^2}$	$\frac{1}{x}$	$e^{-\log^{1/2} x}$	$e^{-\log^{1/4} x}$	$\frac{1}{\log^{\alpha/2} x}$
$e^{-\ell}$	$e^{-\log^2 x}$	$\frac{1}{x}$	$e^{-\log^{1/2} x}$	$\frac{1}{\log^\alpha x}$
$e^{-\ell^{1/2}}$	$e^{-\log^4 x}$	$e^{-\log^2 x}$	$\frac{1}{x}$	$\frac{1}{\log^{2\alpha} x}$
$\frac{1}{\ell^\alpha}$	$e^{-x^{\frac{2}{2+\alpha}}}$	$e^{-x^{\frac{1}{1+\alpha}}}$	$e^{-x^{\frac{1/2}{1/2+\alpha}}}$	$\frac{1}{x}$

Constants omitted  $e^{-c \log^{1/2} x}$ ;  $\frac{1}{x} = e^{-\log x}$

Logarithmic asymptotics

In some corners even log log asymptotics

# Tauberian theorem

$$\mathbb{P}_i(X > x) \sim \int_0^\infty D_i(\ell) \exp\{-\mu\varphi(\ell)x\} f(\ell) d\ell$$

## Theorem

Define  $\bar{\varphi}_I(t) = \int_t^\infty \varphi(y) dy$  and assume

$$f(t) = \varphi(t)\bar{\varphi}_I(t)^{\beta-1} L_0(\bar{\varphi}_I(t))$$

where  $L_0(s)$  is slowly varying at  $s = 0$ . Then

$$\mathbb{P}_i(X > x) \sim D_i^* \frac{\Gamma(\beta)}{\mu^\beta} \frac{L_0(1/x)}{x^\beta}, \quad x \rightarrow \infty.$$

# Rare events approach

$\mathbb{P}(X(\ell) > x)$ : sofar first  $x \rightarrow \infty$ , then  $\ell \rightarrow \infty$ .

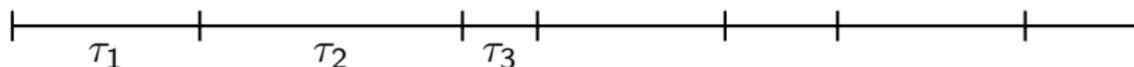
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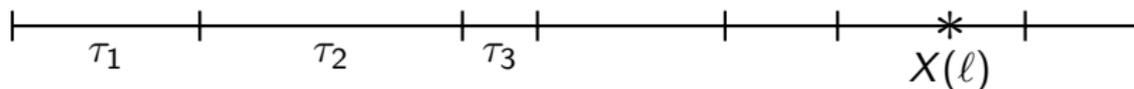
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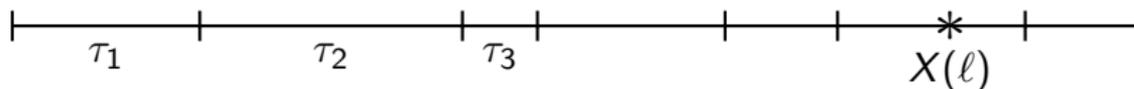
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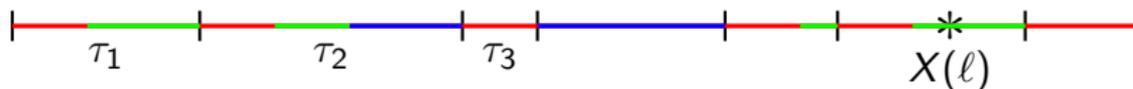
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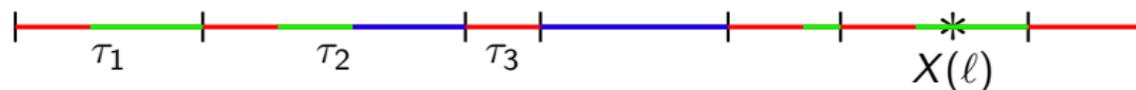
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Express  $\mathbb{E}\tau$ ,  $a(\ell)$  in terms of the  $\pi_i$ ,  $\mathbb{E}_i T$ ,  $\mathbb{P}_u(T_u > \ell)$  etc.

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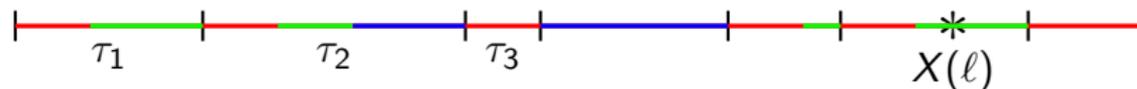
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## Theorem

$$\mathbb{E}X(\ell) \sim \frac{1}{\gamma(\ell)} \sim \frac{1}{\mu\varphi(\ell)}, \quad \ell \rightarrow \infty.$$

# Non-exponential distributions

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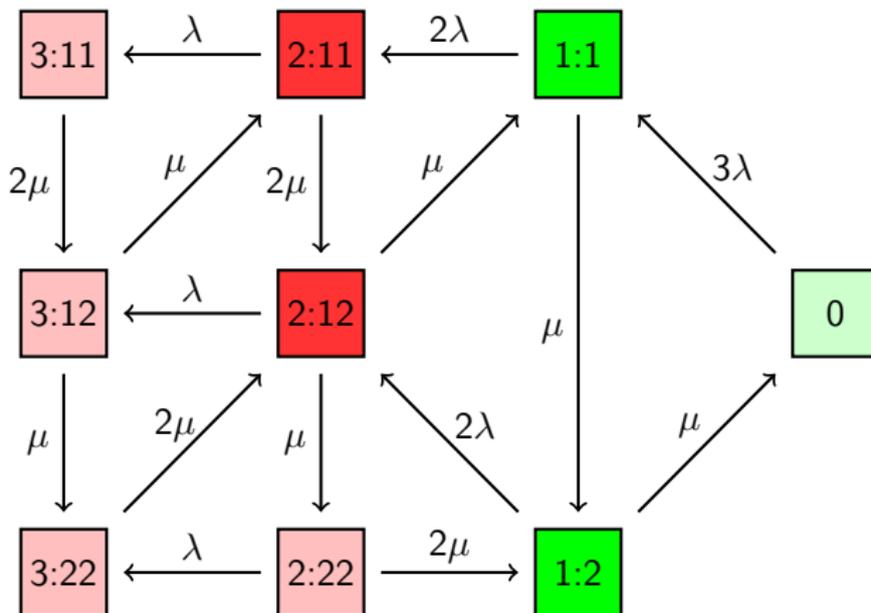
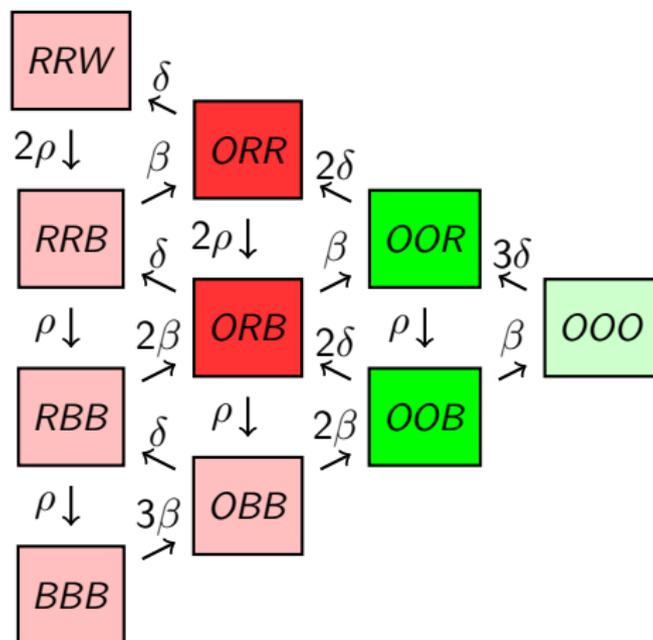


Figure:  $E_2$  repair times, 2-out-of-3

# Variable rates

## Variable rates

2 - out - of - 3 processors, 2 repairmen, warm standby

Rate  $\rho$  of each processorTotal rate  $3\rho$  in OOO,  $2\rho$  in OOO

Task processed at rate  $\rho_u(t)$  in  $u \in \mathcal{U}$

completion at time  $\inf \left\{ s > 0 : \int_0^s \rho_u(t) dt \geq \ell \right\}$  if  $< T_u$

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Formulas in terms of

$$\mathbf{\Delta}_r^{-1} \mathbf{A} - \frac{\delta}{2} (\mathbf{\Delta}_r^{-1} \mathbf{e} \mathbf{e}^T + \mathbf{e} \mathbf{e}^T \mathbf{\Delta}_r^{-1})$$

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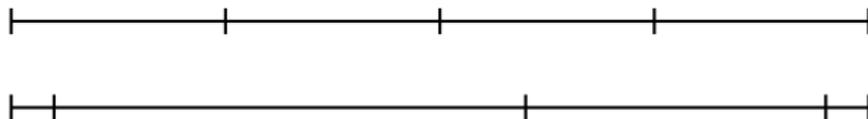
Or  $T_u$  PH

# Fragmentation

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$K$  parts,  $L = L_1 + \dots + L_K$

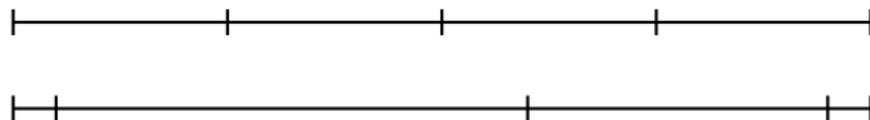
Equidistant, footer and header, etc.



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Equidistant, footer and header, etc.



Parallel computing:

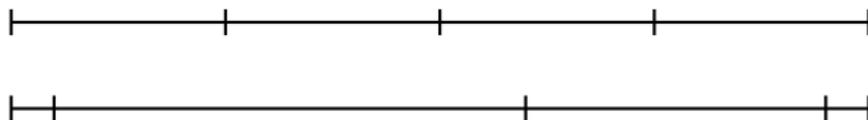
$$X = \max(X_1, \dots, X_K)$$

E.g. Monte Carlo,  $R = R_1 + \dots + R_K$  replications,  $R_k = R/K$

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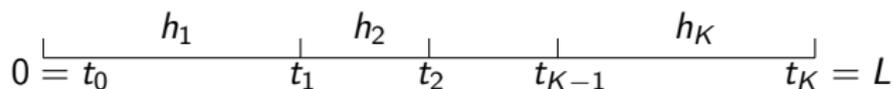
Checkpointing:

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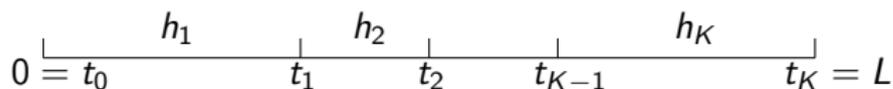
Previous part needs completion

# Checkpoint modeling

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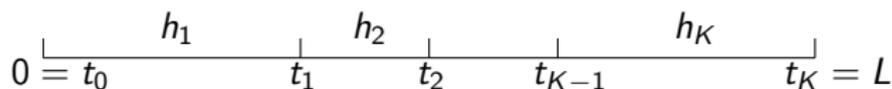
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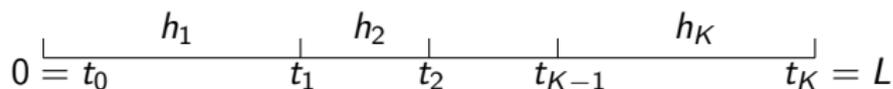
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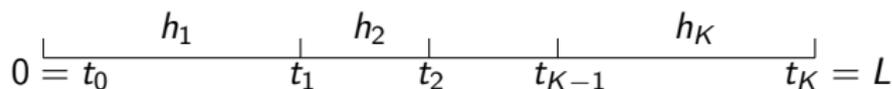
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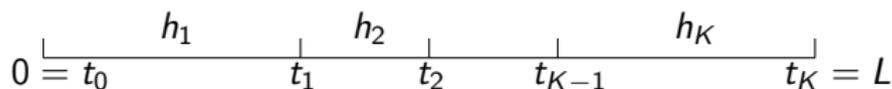
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- B:  $L$  deterministic, checkpoints deterministic but not equally spaced.
- C:  $T$  is deterministic, checkpoints random: outcome of order statistics  $K-1$  i.i.d. uniform r.v.'s on  $(0, t)$ .

# Checkpoint modeling



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 $t_1 = h/K$ ,  $t_2 = 2h/K$ ,  $\dots$ ,  $t_{K-1} = (K-1)h/K$ .
- B:  $L$  deterministic, checkpoints deterministic but not equally spaced.
- C:  $T$  is deterministic, checkpoints random: outcome of order statistics  $K-1$  i.i.d. uniform r.v.'s on  $(0, t)$ .
- D:  $L$  is random and the checkpoints equally spaced,  $h_k \equiv h$ . Thus,  $K = \lceil L/h \rceil$  is random

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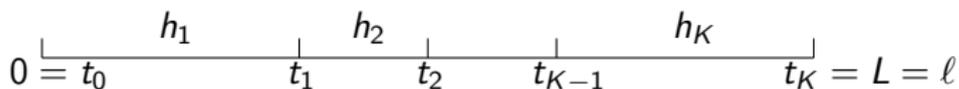
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**Conjecture:** main contribution to  $X$  comes from longest  $L_k$ .  
I.e., it should be best to take  $L_k = L/K$  (Model A).  
  
**NB:** ignores cost of checkpointing  
 $K = \infty$  is optimal if we are free to choose  $K$ ; then  $X = L$

## Which scheme is best given $K$ ?

Take  $L$  deterministic,  $L \equiv \ell$  (Models A,B,C).



- A: Checkpoints are deterministic and equally spaced,  $t_1 = t/K, t_2 = 2t/K, \dots, t_{K-1} = (K-1)t/K$ . Equivalently,  $h_k = t/K$ .
- B: Checkpoints are deterministic but not equally spaced,  $h_k \neq h_\ell$  for  $k \neq \ell$ .
- C: Checkpoints are random: the set  $\{t_1, \dots, t_{K-1}\}$  is the outcome of  $K-1$  i.i.d. uniform r.v.'s on  $(0, t)$ . That is,  $t_1 < \dots < t_{K-1}$  are the order statistics of  $K-1$  i.i.d. uniform r.v.'s on  $(0, t)$ .

**Conjecture:** main contribution to  $X$  comes from longest  $T_k$ .  
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Another possibility:

(ii)  $\mathbb{P}(X_A > x) \leq \mathbb{P}(X_B > x)$  and  $\mathbb{P}(X_A > x) \leq \mathbb{P}(X_C > x)$

for all large  $x$

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**Intuitive explanation:**

Placing checkpoint at 1, failure mechanism starts afresh then.

I.e., the failure rate becomes  $g_1(u) > 0$  instead of  $g_2(u) = 0$ .

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## Theorem

*Assume that the failure rate  $\mu(t) = g(t)/\bar{G}(t)$  of  $G$  is non-decreasing. Then  $X_A(t) \preceq_{st} X_B(t) \preceq_{st} X_C(t) \preceq_{st} X_R(t)$ .*

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**Main case:** Poisson failures.

## Limit theorems models A,B,C.

## Comparison with R

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$$\mathbb{P}(X_R(t) > x) \sim C_R e^{-\gamma(\ell)x},$$

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Assume  $F$  gamma-like,  $f(t) \sim c_F t^\alpha e^{-\lambda t}$ . Then  
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**RESTART comparison:**  $\mathbb{P}(X_R > x) \sim \frac{C_R}{x^{\lambda/\mu}}$

Reduction from power tail to exponential tail.

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**RESTART comparison:**  $\mathbb{P}(X_R > x) \sim C_R \exp\{-\theta \log \log x\}$   
 Heavier than any power; reduction to power tail.

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Assume  $t_k = kT/K$  and that  $F$  is exponential( $\lambda$ ). Then

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Still power tail, but each checkpoint improves the power by 1.

# Thanks

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and to Nick